

Large gaps between random eigenvalues

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Abstract

We show that in the point process limit of the bulk eigenvalues of β -ensembles of random matrices, the probability of having no eigenvalue in a fixed interval of size λ is given by

$$(\kappa_\beta + o(1))\lambda^{\gamma_\beta} \exp\left(-\frac{\beta}{64}\lambda^2 + \left(\frac{\beta}{8} - \frac{1}{4}\right)\lambda\right)$$

as $\lambda \rightarrow \infty$, where

$$\gamma_\beta = \frac{1}{4} \left(\frac{\beta}{2} + \frac{2}{\beta} - 3 \right).$$

This is a slightly corrected version of a prediction by Dyson (1962). Our proof uses the new Brownian carousel representation of the limit process, as well as the Cameron-Martin-Girsanov transformation in stochastic calculus.

1 Introduction

In the 1950s Wigner endeavored to set up a probabilistic model for the repulsion between energy levels in large atomic nuclei. His first models were random meromorphic functions related to random Schrödinger operators, see Wigner (1951) and Wigner (1952). Later, in Wigner (1957) he turned to models of random matrices that are by now standard, such as the Gaussian orthogonal ensemble (GOE). In this model one fills an $n \times n$ matrix M with independent standard normal random variables, then symmetrizes it to get

$$A = \frac{M + M^T}{\sqrt{2}}.$$

The Wigner semicircle law is the limit of the empirical distribution of the eigenvalues of the matrix A . However, Wigner's main interest was the local behavior of the eigenvalues,

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namely the repulsion between them. One mathematical manifestation of this repulsion is large gap probabilities. If we scale the eigenvalue point process of A so that the average spacing is 2π , then Wigner predicted that the probability that there is no eigenvalue in a fixed interval of length λ is given by

$$p_\lambda = \exp \left(-(c + o(1))\lambda^2 \right),$$

where this is a $\lambda \rightarrow \infty$ behavior, and we assume that the $n \rightarrow \infty$ limit has already been taken. This rate of decay is in sharp contrast with the exponential tail for gaps between Poisson points; it is one manifestation of the more organized nature of the random eigenvalues. Wigner's estimate of the constant c , $1/(16\pi)$, later turned out to be inaccurate. Dyson (1962) improved this estimate to

$$p_\lambda = (\kappa_\beta + o(1))\lambda^{\gamma_\beta} \exp \left(-\frac{\beta}{64}\lambda^2 + \left(\frac{\beta}{8} - \frac{1}{4} \right) \lambda \right) \quad (1)$$

where β is a new parameter introduced by noting that the joint eigenvalue density of the GOE is the $\beta = 1$ case of

$$\frac{1}{Z_{n,\beta}} e^{-\beta \sum_{k=1}^n \lambda_k^2/4} \prod_{j < k} |\lambda_j - \lambda_k|^\beta. \quad (2)$$

The family of distributions defined by the density (2) is called the β -ensemble. Dyson's computation of the exponent γ_β , namely $\frac{1}{4}(\frac{\beta}{2} + \frac{2}{\beta} + 6)$, was shown to be slightly incorrect. Indeed, des Cloizeaux and Mehta (1973) gave more substantiated predictions that γ_β is equal to $-1/8$, $-1/4$ and $-1/8$ for values $\beta = 1, 2$ and 4 , respectively. Mathematically precise proofs for the $\beta = 1, 2$ and 4 cases were later given by several authors: Widom (1996), Deift et al. (1997). Moreover, the value of κ_β and higher order asymptotics were also established for these specific cases by Krasovsky (2004), Ehrhardt (2006), Deift et al. (2007). The problem of determining the asymptotic probability of a large gap naturally arises in other random matrix models as well. For a treatment of the case of the β -Laguerre ensemble see Chen and Manning (1996).

Our main theorem gives a mathematically rigorous version of Dyson's prediction for general β with a corrected exponent γ_β .

Theorem 1. *The formula (1) holds with a positive κ_β and*

$$\gamma_\beta = \frac{1}{4} \left(\frac{\beta}{2} + \frac{2}{\beta} - 3 \right).$$

The proof is based on the Brownian carousel, a geometric representation of the $n \rightarrow \infty$ limit of the eigenvalue process. We first introduce the **hyperbolic carousel**. Let

- b be a path in the hyperbolic plane
- z be a point on the boundary of the hyperbolic plane, and
- $f : \mathbb{R}_+ \rightarrow \mathbb{R}_+$ be an integrable function.

To these three objects, the hyperbolic carousel associates a multi-set of points on the real line defined via its counting function $N(\lambda)$ taking values in $\mathbb{Z} \cup \{-\infty, \infty\}$. As time increases from 0 to ∞ , the boundary point z is rotated about the center $b(t)$ at angular speed $\lambda f(t)$. $N(\lambda)$ is defined as the integer-valued total winding number of the point about the moving center of rotation.

The **Brownian carousel** is defined as the hyperbolic carousel driven by hyperbolic Brownian motion b (see Figure 1). It is connected to random matrices via the following theorem:

Theorem 2 (Valkó and Virág (2008)). *Let Λ_n denote the point process given by (2), and let μ_n be a sequence so that $n^{\frac{1}{6}}(2\sqrt{n} - |\mu_n|) \rightarrow \infty$. Then we have the following convergence in distribution:*

$$\sqrt{4n - \mu_n^2}(\Lambda_n - \mu_n) \Rightarrow \text{Sine}_\beta, \quad (3)$$

where Sine_β is the discrete point process given by the Brownian carousel with parameters

$$f(t) = \frac{\beta}{4} e^{-\beta t/4} \quad (4)$$

and arbitrary z .

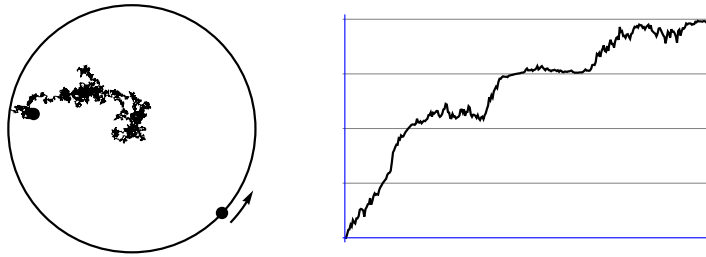


Figure 1: The Brownian carousel and the winding angle α_λ

Remark 3. The semicircle law shows that most points in Λ_n are in the interval $[-2\sqrt{n}, 2\sqrt{n}]$. The discrete point process Λ_n has two kind of point process limits, one near the edges of this interval and another in the bulk. The condition on the parameter μ_n means that we get a bulk-type scaling limit of Λ_n . The scaling factor in (3) is the natural choice in view of the Wigner semicircle law in order to get a point process with average density $1/(2\pi)$. The limiting point process for the edge-scaling case have been obtained by Ramirez, Rider, and Virág (2007).

The Brownian carousel description gives a simple way to analyze the limiting point process. The hyperbolic angle of the rotating boundary point as measured from $b(t)$ follows the following coupled one-parameter family of stochastic differential equations

$$d\alpha_\lambda = \lambda f dt + \operatorname{Re}((e^{-i\alpha_\lambda} - 1)dZ), \quad \alpha_\lambda(0) = 0, \quad (5)$$

driven by a two-dimensional standard Brownian motion and f given in (4). For a single λ , this reduces to the one-dimensional stochastic differential equation

$$d\alpha_\lambda = \lambda f dt + 2 \sin(\alpha_\lambda/2) dW, \quad \alpha_\lambda(0) = 0, \quad (6)$$

which converges as $t \rightarrow \infty$ to an integer multiple $\alpha_\lambda(\infty)$ of 2π . In particular, the number of points of the point process Sine_β in $[0, \lambda]$ has the same distribution as $\alpha_\lambda(\infty)/(2\pi)$ and p_λ is equal to the probability that α converges to 0 as $t \rightarrow \infty$. See Valkó and Virág (2008) for further details.

In the analysis of equation (6) it helps to remove the space dependence from the diffusion coefficient by a change of variables $X(t) = \log(\tan(\alpha(t)/4))$. The diffusion X satisfies the stochastic differential equation:

$$dX = \frac{\lambda}{2} f \cosh X dt + \frac{1}{2} \tanh X dt + dB, \quad X(0) = -\infty. \quad (7)$$

In Valkó and Virág (2008) equations (6) and (7) were used to identify the leading term in the asymptotic expansion of p_λ in (1). The proof of Theorem 1 requires a more careful analysis of equation (7).

In Lemma 4 we will show that for any initial condition $X(0) = x \in [-\infty, \infty)$ there is a unique solution of the equation given in (7) and $p_\lambda = p_\lambda(-\infty)$ where

$$p_\lambda(x) := \mathbf{P}(X(t) \text{ is finite for all } t > 0 \text{ and does not converge to } \infty \text{ as } t \rightarrow \infty). \quad (8)$$

A time shift of equation (7) only changes the parameter λ and the initial condition. This, together with the Markov property of the diffusion X , shows that with $T = \frac{4}{\beta} \log \lambda$ we

have

$$p_\lambda = \mathbf{E} \left[\mathbf{1} \{X(t) \text{ is finite for all } 0 < t \leq T\} \cdot p_1(X(T)) \right]. \quad (9)$$

Our main tool is the Cameron-Martin-Girsanov formula, which allows one to compare the measure on paths given by two diffusions. If we knew the conditional distribution of the diffusion X under the event that it does not blow up, then we could use the Cameron-Martin-Girsanov formula to compute p_λ explicitly. While we cannot do this, the next best option is to find a new diffusion Y which approximates this conditional distribution. The density (i.e. the Radon-Nikodym derivative) of the path measures given by Y with respect to the measure given by X will be close to the right hand side of (1). Our strategy for finding Y is described in Section 4.

In Section 5, we will present a coupling of the transformed processes that enables us to show that the asymptotics is precise up to and including the constant term κ_β . The term κ_β is then identified as the expectation of a functional of a certain limiting diffusion.

Open problem 1. Give an explicit expression for κ_β for general values of β .

The known values of κ_β are

$$\kappa_1 = 2^{13/24} e^{\frac{3}{2}\zeta'(-1)}, \quad \kappa_2 = 2^{7/12} e^{3\zeta'(-1)}, \quad \kappa_4 = 2^{-13/12} e^{\frac{3}{2}\zeta'(-1)},$$

where $\zeta'(-1)$ is the coefficient of the linear term in the Laurent series of the Riemann- ζ function at -1.

A natural generalization of Theorem 1 would be to consider the asymptotic probability that there are exactly n eigenvalues in a large interval $[0, \lambda]$. This probability is usually denoted by $E_\beta(n; \lambda)$ in the literature. For $\beta = 1, 2$ and 4 the following large λ asymptotics was obtained by Basor et al. (1992):

$$\log E_\beta(n; \lambda) = \log E_\beta(0; \lambda) + \frac{n\beta}{4}\lambda + \frac{n}{2} \left(1 - \frac{\beta}{2} - \frac{n\beta}{2} \right) \log \lambda + c_\beta + o(1)$$

with explicit constants c_β . (See also Tracy and Widom (1993).) We believe that our methods can be used to extend the previous asymptotics for general values of β .

The rest of the paper is organized as follows. In the next section, we justify the $p_\lambda = p_\lambda(-\infty)$ and give some preliminary estimates on the probability $p_1(r)$ appearing in (8). Section 3 presents the version of the Cameron-Martin-Girsanov formula that we need. Section 4 describes the strategy for finding Y and Section 5 builds on these sections to complete the proof of the main theorem.

2 Preliminary results

First we formally verify the connection between the gap probability p_λ and the diffusion given in (7).

Lemma 4. *The diffusion (7) has a unique solution for any initial condition $X(0) = x \in [-\infty, \infty)$ and $p_\lambda = p_\lambda(-\infty)$.*

Proof. The change of variables function $\log \tan(\cdot/4)$ is one-to-one on $(0, 2\pi) \rightarrow \mathbb{R}$. Therefore, even with $-\infty$ initial condition, the diffusion X is well defined and has a unique solution until α reaches 2π , when it blows up. We define $X(t) = \infty$ after this blowup.

Note that for $\lambda > 0$ the solution of equation (6) is always monotone increasing at multiples of 2π . See Section 2.2 in Valkó and Virág (2008) for more details. So if $\alpha(t) \rightarrow 0$ as $t \rightarrow \infty$ then $0 < \alpha(t) < 2\pi$ for all $t > 0$. This means that $X(t)$ is finite for all $t > 0$ and $X(t)$ cannot converge to ∞ which proves $p_\lambda = p_\lambda(-\infty)$. \square

Next, we prove a preliminary estimate on the blowup probability of the diffusion (7).

Lemma 5. *Recall that $p_1(x)$ is the probability that the diffusion (7) with $\lambda = 1$ and initial condition $X(0) = x$ does not blow up in finite time and does not converge to $+\infty$ as $t \rightarrow \infty$. We have*

$$0 < p_1(x) \leq c_\beta \exp\left(-\frac{\beta}{60} e^x\right).$$

Proof. For the upper bound we first assume that $x > 4$. Consider the diffusion

$$dR = \frac{\beta}{16} e^{R-\beta/4 t} dt + dB, \quad R(0) = x. \quad (10)$$

This has the same noise term as X . The drift term of R is less than the drift term of X (with $\lambda = 1$) as functions depending on time and space when the space variables are nonnegative. Thus while R stays positive, we have $R \leq X$. This means that for every $t > 0$ we have

$$\begin{aligned} p_1(x) &\leq \mathbf{P}(X \text{ does not blow up before time } t) \\ &\leq \mathbf{P}\left(\min_{s \in [0, t]} R(s) < 0 \text{ or } R \text{ does not blow up before time } t\right). \end{aligned} \quad (11)$$

The difference $Z = R - B$ satisfies the ODE

$$e^{-Z} dZ = \frac{\beta}{16} e^{B-\beta/4 t} dt, \quad Z(0) = x. \quad (12)$$

Integration gives

$$e^{-x} - e^{-Z(t)} = \frac{\beta}{16} \int_0^t e^{B(s) - \beta/4 s} ds.$$

This shows that Z is increasing in t , in particular $Z(t) \geq x$. So if $\min_{[0,t]} R < 0$ then

$$\min_{[0,t]} B < -x.$$

Furthermore, if

$$e^{-x} < \frac{\beta}{16} \int_0^t e^{B(s) - \beta/4 s} ds \quad (13)$$

then R blows up before time t . This certainly happens if the minimum of B on the interval $[0, t]$ is not sufficiently small. More precisely, if

$$\frac{e^{-b}}{4}(1 - e^{\beta t/4}) > e^{-x} \quad (14)$$

and $\min_{[0,t]} B > -b$ then (13) follows. So if $b < x$, and (14) holds, then the right hand side of (11) can be bounded above by

$$P\left(\min_{[0,t]} B < -b\right) = P(|B(t)| > b) \leq \frac{\sqrt{t}}{b} e^{-\frac{b^2}{2t}}.$$

We set

$$t = \frac{16}{\beta} e^{2-x}, \quad b = \frac{4e}{\sqrt{30}} < 2.$$

As $x > 4$, both $b < x$ and (14) are satisfied and we get the upper bound

$$p_1(x) \leq \frac{\sqrt{t}}{b} e^{-\frac{b^2}{2t}} < c_\beta e^{-\frac{\beta}{60} e^x}$$

with $c_\beta = \sqrt{30/\beta}$. The upper bound for all values of x now follows by changing the constant c_β appropriately.

For the lower bound note that since the Sine_β process is discrete and translation invariant in distribution, there exists $\nu \in (0, 1)$ so that $p_\nu = p_\nu(-\infty) > 0$. By the Markov property, we have

$$p_\nu = \int_{-\infty}^{\infty} K_{0,1}(-\infty, dx) p_{\nu e^{\beta/4}}(x)$$

where $K_{s,t}(y, dx)$ is the transition kernel of the Markov process X with parameter $\lambda = \nu$. This implies that for some $x_0 \in \mathbb{R}$ we have

$$p_{\nu e^{\beta/4}}(x_0) > 0.$$

Consider the process X started at x with parameter $\lambda = 1$. The Markov property applied at time $t_0 = 1 - \frac{4}{\beta} \log \nu$ and the monotonicity of $p_\lambda(x)$ in x implies

$$p_1(x) \geq P(X(t_0) < x_0) p_{\nu e^{\beta/4}}(x_0) > 0,$$

since $P(X(t) < x)$ is positive for all $x \in \mathbb{R}$ and $t > 0$. □

3 The Cameron-Martin-Girsanov formula

The diffusions we work with can blow up to $+\infty$ in finite time, in which case they are required to stay there forever after. For this reason, we will not be able to use the textbook versions of the Cameron-Martin-Girsanov formula, and have to prove a version that applies to such diffusions.

Proposition 6. *Consider the following stochastic differential equations*

$$dX = g(t, X)dt + dB, \quad \lim_{t \rightarrow 0} X(t) = -\infty \quad (15)$$

$$dY = h(t, Y)dt + d\tilde{B}, \quad \lim_{t \rightarrow 0} Y(t) = -\infty \quad (16)$$

on the interval $(0, T]$, and assume that (15) has a unique solution X in law taking values in $(-\infty, \infty]$.

Let

$$G_s = G_s(X) = \int_0^s h(t, X) - g(t, X)dX - \frac{1}{2} \int_0^s h(t, X)^2 - g(t, X)^2 dt \quad (17)$$

and assume that

- (A) $g^2 - h^2$ and $g - h$ are bounded when x is bounded above. (Then G_s is almost surely well-defined when X_s is finite.)
- (B) G_s is bounded above by a deterministic constant.
- (C) $G_s \rightarrow -\infty$ when $s \uparrow \tau$ if X hits $+\infty$ at time τ . In this case we define $G_s := -\infty$ for $s \geq \tau$.

Consider the process \tilde{Y} whose density with respect to the distribution of the process X is given by e^{G_T} . Then \tilde{Y} satisfies the second SDE (16) and never blows up to $+\infty$ almost surely. Moreover, for any nonnegative function φ of the path of X that vanishes when X blows up we have

$$\mathbf{E} \varphi(X) = \mathbf{E} [\varphi(Y)e^{-G_T(Y)}]. \quad (18)$$

Proof. We follow the standard proof of the Girsanov theorem.

First we show that G_s is well defined for finite X_s . From condition (A) it follows that the second integral is well-defined. The first integral can be written as

$$\int_0^s (h - g)dB + \int_0^s (h - g)g dt$$

which is well defined since $(h - g)$ and $2(h - g)g = (h^2 - g^2) - (h - g)^2$ when their argument x is bounded above.

Next, we show that $M_s = e^{G_s}$ is a bounded martingale. This is clear after the hitting time τ of X of $+\infty$, if such time exists. Before this time, G_s is a semimartingale, and so is M_s . Itô's formula gives

$$dM = (h - g)MdB$$

so that the drift term of M vanishes. So M is a local martingale which is bounded, so it has to be a martingale.

The rest of the proof is standard, and follows the following outline. Set

$$\tilde{B}_s = X_s - \int_0^s h dt = B_s - \int_0^s (h - g)dt.$$

It suffices to show that \tilde{B} is a Brownian motion with respect to the new measure with density M_T . This follows from Lévy's criterion (Karatzas and Shreve (1991) Theorem 3.3.16) if \tilde{B} and $\tilde{B}^2 - s$ are local martingales. Since M is a martingale, it suffices to show that $\tilde{B}M$ and $(\tilde{B}^2 - s)M$ are local martingales with respect to the old measure, which is just a simple application of Itô's formula.

The identity (18) is just a version of the change of density formula. □

4 Construction of the diffusion Y

In this section we will create a diffusion which approximates the conditional distribution of the diffusion X under the event that it does not blow up. We will construct a drift function $h(t, x)$ for which the diffusion Y

$$dY = h(t, Y)dt + dB_t, \quad Y(0) = -\infty \tag{19}$$

is well-defined, a.s. finite for $t > 0$ and the (formal) Radon-Nikodym derivative e^{G_T} with G_T defined in (17) is almost equal to the right hand side of equation (1) with the appropriate γ_β .

Lemma 7. *For the diffusion (7), $\lambda > 1$ and $T = \frac{4}{\beta} \log \lambda$ there exists a function $h(t, x)$ so that conditions (A)-(C) of Proposition 6 hold, and G_T has the following form:*

$$\begin{aligned} -G_T(X) &= -\frac{\beta}{64} \lambda^2 + \left(\frac{\beta}{8} - \frac{1}{4}\right) \lambda + \frac{1}{8} \left(\beta + \frac{4}{\beta} - 6\right) \log \lambda \\ &\quad + \frac{\beta}{8} e^{X(T)} + \left(2 - \frac{\beta}{2}\right) X(T)^+ + \omega(X(T)) + \int_0^T \phi(T - t, X(t))dt. \end{aligned} \tag{20}$$

Here the function ω is bounded and continuous, ϕ is continuous and $|\phi(t, x)| \leq \tilde{\phi}(t)$ with $\int_0^\infty \tilde{\phi}(t)dt < \infty$. The functions ω and $\phi, \tilde{\phi}$ may depend on the parameter β , but not on λ .

The function h will have the following form

$$h(t, x) = -\frac{\lambda}{2}f \sinh(x) + h_0(t, x), \quad (21)$$

where $|h_0(t, x)| < c$ if $0 \leq t \leq T$. The constant c depends only on β .

Proof. Construction of the function h . Given an explicit formula for h it would not be hard to check that G_T has the desired form. However, we would like to present a way one can find the appropriate drift function. This will provide a better understanding of the form of the resulting h .

We will use the definition

$$-G_s(X) = \int_0^s g(t, X) - h(t, X) dX + \frac{1}{2} \int_0^s h^2(t, X) - g^2(t, X) dt.$$

where

$$g = g_1 + g_2, \quad g_1(t, x) = \frac{\lambda}{2}f(t) \cosh x, \quad g_2(t, x) = \frac{1}{2} \tanh x.$$

Our goal is to find the appropriate drift term h in a way that the diffusion Y will approximate the conditional distribution of X given that it does not blow up in the interval $[0, T]$. We will do this term by term, starting with the highest order; towards this end we write $h = h_1 + h_2 + h_3 + h_4$. We set

$$h_1(t, x) = -\frac{\lambda}{2}f(t) \sinh(x) \quad (22)$$

as this yields the nice cancellation

$$h_1^2 - g_1^2 = \frac{\lambda^2}{4}f(t)^2 \sinh^2(X) - \frac{\lambda^2}{4}f(t)^2 \cosh^2(X) = -\frac{\lambda^2}{4}f(t)^2$$

in the main terms of $h^2 - g^2$. In addition, if the remaining term $h_2 + h_3 + h_4$ is bounded, then it will be easy to show that conditions (A)-(C) of Proposition 6 are satisfied. This will be done at the end of the proof.

The contribution of the drift terms h_1 and g_1 to the stochastic integral part of $-G_s$ is given by

$$\frac{\lambda}{2} \int_0^s f(t)(\cosh(X) - \sinh(X)) dX = \frac{\lambda}{2} \int_0^s e^X f(t) dX. \quad (23)$$

Our main tool for evaluating integrals with respect to dX is the following version of Itô's formula. Let a, b be continuously differentiable functions and let \tilde{a} denote the antiderivative of a . Then

$$a(t)b(X) dX = d(a(t)\tilde{b}(X)) - a'(t)\tilde{b}(X) dt - \frac{1}{2}a(t)b'(X) dt. \quad (24)$$

Since $f'(t) = -\beta/4f(t)$, and $X(0) = -\infty$, this formula gives

$$\frac{\lambda}{2} \int_0^s f(t) e^X dX = \frac{\lambda}{2} f(s) e^{X(s)} + \frac{\lambda}{2} \left(\frac{\beta}{4} - \frac{1}{2} \right) \int_0^s e^X f dt. \quad (25)$$

Next we would like to choose h_2 in (22) so that the integral term in the right hand side of (25) simplifies. More precisely, since we expect the diffusion X to be near 0 most of the time, we would like to replace the term e^X by 1. The plan is to use the cross term $\int h_1 h_2 dt$ in the $\frac{1}{2} \int h^2 dt$ term of G to do this. Namely, we would like to have

$$h_1 h_2 = \frac{\lambda}{2} \left(\frac{\beta}{4} - \frac{1}{2} \right) (1 - e^x) f. \quad (26)$$

The solution for (26) is given by

$$h_2(t, x) = \left(\frac{\beta}{4} - \frac{1}{2} \right) (1 + \tanh(x/2)). \quad (27)$$

We will choose the next term, h_3 , so that the cross term $\int h_1 h_3 dt$ in $\frac{1}{2} \int h^2 dt$ cancels the cross term $-\int g_1 g_2$ in $-\frac{1}{2} \int g^2 dt$. This leads to the equation

$$h_1 h_3 = g_1 g_2 = \frac{\lambda}{2} f(t) \cosh(x) \cdot \frac{1}{2} \tanh(x),$$

which gives

$$h_3(t, x) = -\frac{1}{2}. \quad (28)$$

Collecting all our previous computations we get

$$\begin{aligned} -G_s &= \frac{\lambda}{2} f(s) e^{X(s)} - \frac{\lambda^2}{8} \int_0^s f^2 dt + \lambda \left(\frac{\beta}{8} - \frac{1}{4} \right) \int_0^s f dt \\ &\quad + \frac{1}{2} \int_0^s 2h_1 h_4 + (h_2 + h_3 + h_4)^2 - g_2^2 dt \\ &\quad - \int_0^s h_4 dX + \int_0^s g_2 - h_2 - h_3 dX. \end{aligned} \quad (29)$$

The integrand $u = g_2 - h_2 - h_3$ in the last integral of (29) has antiderivative

$$\tilde{u}(x) = \left(1 - \frac{\beta}{4} \right) x + \left(1 - \frac{\beta}{2} \right) \log \cosh(x/2) + \frac{1}{2} \log \cosh x. \quad (30)$$

By Itô's formula, $\int_0^s u(X) dX - \tilde{u}(X)|_0^s$ is given by

$$\begin{aligned} -\frac{1}{2} \int_0^s u'(X) dt &= -\frac{1}{2} \int_0^s \left[\frac{2-\beta}{8} \operatorname{sech}(X/2)^2 + \frac{1}{2} \operatorname{sech}(X)^2 \right] dt \\ &= \frac{\beta-6}{16} s + \int_0^s \left[\frac{(2-\beta)}{16} \tanh(X/2)^2 + \frac{1}{4} \tanh(X)^2 \right] dt. \end{aligned} \quad (31)$$

Note that

$$\lim_{x \rightarrow -\infty} \tilde{u}(x) = \frac{\beta - 3}{2} \log 2 = c_1.$$

Substituting this computation for the last integral and expanding $(h_2 + h_3 + h_4)^2$ we can rewrite (29) as follows.

$$\begin{aligned} -G_s &= -\frac{\lambda^2}{8} \int_0^s f^2 dt + \lambda \left(\frac{\beta}{8} - \frac{1}{4} \right) \int_0^s f dt + \left(\frac{1}{2} \left(\frac{\beta}{4} - 1 \right)^2 + \frac{\beta - 6}{16} \right) s \\ &\quad + \frac{\lambda}{2} f(s) e^{X(s)} + \tilde{u}(X(s)) - c_1 \\ &\quad + \frac{1}{2} \int_0^s (2h_1 h_4 + 2(h_2 + h_3)h_4 + h_4^2) dt - \int_0^s h_4 dX + \int_0^s \eta(X(t)) dt. \end{aligned} \quad (32)$$

The coefficient of s in the first line of (32) comes from the first term on the right in (31) and the constant term of $(h_2 + h_3)^2/2$. The function η collects the terms from the integrand in (31), the terms $(h_2 + h_3)^2/2$ with the constant term $(\beta/4 - 1)^2/2$ removed, and $-g_2^2/2$. More explicitly, we have

$$\eta(x) = \frac{(8 - 6\beta + \beta^2)}{32} (2 \tanh(x/2) + \tanh(x/2)^2) + \frac{1}{8} (\tanh x)^2.$$

The function $\eta(x)$ contributes to an error term that needs to be controlled, but whose precise value does not influence our final result. Now we are ready to set the value for h_4 : we will choose it in a way that the cross term $\int h_1 h_4 dt$ in (32) will cancel the integral $\int \eta dt$. This gives $h_4 = -\eta/h_1$, that is

$$h_4 = \frac{2}{\lambda f(t)} \frac{\eta(x)}{\sinh(x)}. \quad (33)$$

The function h_4 is a product of a function of t and a function of x . Itô's formula (24), with the notation $\tilde{h}_4(t, x) = \int_0^x h_4(t, y) dy$ yields the evaluation of the stochastic integral in (32):

$$-\int_0^s h_4 dX = -\tilde{h}_4(s, X(s)) + \frac{\beta}{4} \int_0^s \tilde{h}_4 dt + \frac{1}{2} \int_0^s \partial_x h_4 dt.$$

Plugging this into (32) and simplifying the deterministic terms in the first line of (32) we arrive at

$$\begin{aligned} -G_s &= -\frac{\lambda^2}{8} \int_0^s f^2 dt + \lambda \left(\frac{\beta}{8} - \frac{1}{4} \right) \int_0^s f dt + \frac{1}{32} (\beta^2 + 12\beta + 8) s \\ &\quad + \lambda e^{-\beta/4s} \frac{\beta}{8} e^{X(s)} + \tilde{u}(X(s)) - c_1 - \tilde{h}_4(s, X(s)) \\ &\quad + \int_0^s \left(2(h_2 + h_3)h_4 + h_4^2 + \frac{\beta}{4} \tilde{h}_4 + \frac{1}{2} \partial_x h_4 \right) dt. \end{aligned} \quad (34)$$

Note that h_2 and h_3 do not depend on t and are bounded by an absolute constant. The functions $h_4, \tilde{h}_4, \partial_x h_4$ are all bounded by a constant times $1/(\lambda f(t)) = \frac{16}{\beta^2} f(T-t)$, which itself is bounded by a constant not depending on λ as long as $0 \leq t \leq T$. Thus we can rewrite the integrand in (34) as

$$\int_0^s \phi(T-t, X(t)) dt \quad (35)$$

with a continuous function ϕ which does not depend on λ and satisfies $|\phi(t, x)| \leq \tilde{\phi}(t)$ with $\int_0^\infty \tilde{\phi}(t) dt < \infty$. Using (30) and the fact that $\log \cosh x - |x|$ is bounded, the terms in the second line of (34) can be written as

$$\left(2 - \frac{\beta}{2}\right) X(s)^+ + \lambda e^{-\beta/4s} \frac{\beta}{8} e^{X(s)} + \omega_0(X(s)) - \tilde{h}_4(T, X(T)) \quad (36)$$

with a bounded and continuous ω_0 . This concludes the construction of the function h . In order to get the expression (20) for $-G_T$ we first plug in $s = T$ into (34). Then the first line gives

$$-\frac{\lambda^2 \beta}{64} (1 - \lambda^{-2}) + \lambda \left(\frac{\beta}{8} - \frac{1}{4} \right) (1 - \lambda^{-1}) + \frac{1}{4} \left(\frac{\beta}{2} + \frac{2}{\beta} - 3 \right) \log \lambda,$$

and by (36) the second line transforms to

$$\left(2 - \frac{\beta}{2}\right) X(s)^+ + \frac{\beta}{8} e^{X(s)} + \omega_0(X(T)) - \tilde{h}_4(T, X(T)).$$

Note that the expression $\tilde{h}_4(T, x)$ does not depend on T and is bounded. This proves that $-G_T$ is in the desired form (20).

Now we are ready to check that the proposed choice of h satisfies all the needed conditions (A)-(C).

Condition (A). As $x \rightarrow -\infty$ we have

$$g(t, x) = \frac{1}{4} \lambda f e^{-x} - \frac{1}{2} + \hat{g}(t, x), \quad h(t, x) = \frac{1}{4} \lambda f e^{-x} - \frac{1}{2} + \hat{h}(t, x)$$

where $|\hat{g}| < c e^x$ and $|\hat{h}| < c' e^x$ with constants that only depend on β if $0 \leq t \leq T$. From this it follows that $g - h$ and $g^2 - h^2$ are both bounded if x is bounded from above.

Condition (B). We need that (34) is bounded from below if $0 \leq s \leq T$. The integrals in the first line are bounded by a constant depending on λ and β only. The same is true for the integral in the last line, see (35) and the discussion around it. Thus we only need to deal with the evaluation terms of the second line. By (36) we just need to show that

$$\left(2 - \frac{\beta}{2}\right) X(s)^+ + \lambda e^{-\beta/4s} \frac{\beta}{8} e^{X(s)} \quad (37)$$

is bounded from below. Since $s \leq T = \frac{4}{\beta} \log \lambda$, we get that (37) is bounded from below by

$$\left(2 - \frac{\beta}{2}\right) X(s)^+ + \frac{\beta}{8} e^{X(s)}$$

which in turn is bounded from below by a constant depending only on β .

Condition (C). This follows the same way: one only needs to check the behavior of (37) as s converges to the hitting time of ∞ . This expression converges to ∞ as $X(s) \rightarrow \infty$ which means that $G_s \rightarrow -\infty$. \square

5 The proof of the main theorem

We are ready to prove Theorem 1.

Proof of Theorem 1. Lemma 4 gives $p_\lambda = p_\lambda(-\infty)$, where

$$p_\lambda(x) = \mathbf{P}(X(t) \text{ is finite for all } t > 0 \text{ and does not go to } \infty \text{ as } t \rightarrow \infty)$$

with $X(0) = x$, as defined in (8). Note that a time shift of equation (7) only changes λ and the initial condition. With

$$T = T_\lambda = \frac{4}{\beta} \log \lambda \quad (38)$$

the diffusion $\tau \mapsto X(\tau + T)$ satisfies (7) with $\lambda = 1$ and with initial condition $-\infty$ at $\tau = -T$. This suggests that the dependence on λ for the probability on the right hand side of (8) comes mainly from the interval $[0, T]$. Because of this we take conditional expectations in (8) with respect to the σ -algebra generated by $(X(t), t \in [0, T])$. Using the Markov property of X we obtain

$$p_\lambda = \mathbf{E}(1 \{X(t) \text{ is finite for all } 0 < t \leq T\} \cdot p_1(X(T))) \quad (39)$$

The first term in the expectation is a function of the path $X(t)$ on the time interval $[0, T]$. Consider a diffusion Y given by the SDE (19) with a drift function $h(t, x)$ given by Lemma 7. With the notation of Lemma 7 we set

$$\psi(Y) = \left(2 - \frac{\beta}{2}\right) Y(T)^+ + \frac{\beta}{8} e^{Y(T)} + \omega(Y(T)) + \int_0^T \phi(T - t, Y(t)) dt. \quad (40)$$

We apply the Girsanov transformation of Proposition 6 together with equation (20) of Lemma 7 to get

$$p_\lambda = \lambda^{\gamma_\beta} e^{-\frac{\beta}{64} \lambda^2 + (\frac{\beta}{8} - \frac{1}{4}) \lambda} \mathbf{E}[p_1(Y(T_\lambda)) \exp\{\psi(Y)\}],$$

where $\gamma_\beta = \frac{1}{4} \left(\frac{\beta}{2} + \frac{2}{\beta} - 3 \right)$. In order to prove the theorem it suffices to show that the limit

$$\lim_{\lambda \rightarrow \infty} \mathbf{E} [p_1(Y(T_\lambda)) \exp\{\psi(Y)\}] \quad (41)$$

exists, and is finite and positive. This limit then equals the constant κ_β of the asymptotics. Recall that in (40) the function ω is continuous and bounded and $\phi(t, y)$ can be dominated by a function $\tilde{\phi}(y)$ which has a finite integral in $[0, \infty)$.

We will run the process $Y_\lambda(t)$ with a shifted time, $\tau = t - T = t - \frac{4}{\beta} \log \lambda$; that is, let

$$\tilde{Y}_T(\tau) := Y_\lambda(\tau + T).$$

The advantage of this shifted time is that the diffusions $\tilde{Y}_T(\tau)$ for different λ satisfy the same SDE except they evolve on nested time intervals:

$$d\tilde{Y}_T(\tau) = \tilde{h}(\tau, \tilde{Y}) d\tau + dB, \quad \tau > -T, \quad \tilde{Y}_T(-T) = -\infty, \quad (42)$$

where the drift term is given by

$$\tilde{h}(\tau, y) = h(T + \tau, y) = -\frac{\beta}{8} e^{-\beta\tau/4} \sinh(y) + h_0(T + \tau, y). \quad (43)$$

In this new time-frame we need to show that the limit

$$\lim_{T \rightarrow \infty} \mathbf{E} p_1(\tilde{Y}_T(0)) \exp\{\tilde{\psi}(\tilde{Y}_T)\} \quad (44)$$

exists, is positive and finite, where

$$\tilde{\psi}(\tilde{Y}) = \left(2 - \frac{\beta}{2}\right) \tilde{Y}(0)^+ + \frac{\beta}{8} e^{\tilde{Y}(0)} + \omega(\tilde{Y}(0)) + \int_0^T \phi(t, \tilde{Y}(-t)) dt. \quad (45)$$

We will drive the diffusions (42) with the same Brownian motion $B(t)$. Then for $T_1 > T_2$ we have $Y_{T_1}(\tau) > Y_{T_2}(\tau)$ for $\tau \in [T_2, \infty)$ as this holds for $\tau = -T_2$ and the domination is preserved by the evolution.

We also consider a nonnegative-valued diffusion $Z(t)$ given by the SDE

$$dZ = r(Z)dt + dB$$

which is reflected at 0 and whose drift term is equal to

$$r(y) = -\frac{\beta}{16} e^y + c_1. \quad (46)$$

By Lemma 7 the term $h_0(y, T + \tau)$ in (43) is bounded if $-T \leq \tau \leq 0$. Thus we can choose the constant c_1 in (46) so that

$$r(z) \geq \sup_{\{\tau < 0, 0 \leq y \leq z\}} h(\tau, y). \quad (47)$$

Since Z and \tilde{Y} are driven by the same Brownian motion, if $Z, \tilde{Y} > 0$ then $Z - \tilde{Y}$ evolves according to

$$d(Z - \tilde{Y}) = [r(Z) - f(t, Y)] dt.$$

By (47) this means that if $Z(\tau_0) \geq \tilde{Y}(\tau_0)$ for a $\tau_0 < 0$ then this ordering is preserved by the coupling until time 0.

Consider the process Z in its stationary distribution. Then $Z(-T) \geq \tilde{Y}_T(-T)$ therefore Z dominates \tilde{Y}_T on $[-T, 0]$. For every fixed $\tau \leq 0$ the random variables $\tilde{Y}_T(\tau)$ are increasing in T and bounded by $Z(\tau)$ so

$$\tilde{Y}_\infty(\tau) = \lim_{T \rightarrow \infty} \tilde{Y}_T(\tau)$$

exists and is dominated by $Z(\tau)$. The function $p_1(x)$ is continuous so $p_1(\tilde{Y}_T) \rightarrow p_1(\tilde{Y}_\infty)$. By (45) we have

$$\tilde{\psi}(\tilde{Y}_T) = a(\tilde{Y}_T(0)) + \int_0^T \phi(t, \tilde{Y}(-t)) dt$$

where a is continuous and $\phi(t, y)$ can be dominated by a function $\tilde{\phi}(y)$ which has a finite integral in $[0, \infty)$. Hence $\tilde{\psi}(\tilde{Y}_T) \rightarrow \tilde{\psi}(\tilde{Y}_\infty)$ and

$$q_T = e^{\tilde{\psi}(\tilde{Y}_T)} p_1(\tilde{Y}_T) \rightarrow q_\infty = e^{\tilde{\psi}(\tilde{Y}_\infty)} p_1(\tilde{Y}_\infty), \quad \text{as } T \rightarrow \infty.$$

Using Lemma 5 to estimate $p_1(y)$ we get

$$q_T \leq c \exp \left\{ (2 - \beta/2) \tilde{Y}_T(0)^+ + \frac{\beta}{8} e^{\tilde{Y}_T(0)} - \frac{\beta}{60} e^{\tilde{Y}_T(0)} \right\} \leq c' \chi(\tilde{Y}_T(0)),$$

where $\chi(y) = \exp \left\{ \left(\frac{\beta}{8} - \frac{\beta}{60} \right) e^y \right\}$. If we prove that $\mathbf{E} \chi(Z(0)) < \infty$ then the dominated convergence theorem will imply

$$\mathbf{E} q_T \rightarrow \mathbf{E} q_\infty < \infty, \tag{48}$$

and the existence of the limiting constant κ_β will be established.

The generator of the reflected diffusion Z is given by

$$\mathcal{L}f = \frac{1}{2} f'' + f' r,$$

for functions f defined on $[0, \infty)$ with $f'(0+) = 0$. (Revuz and Yor (1999), Chapter VII. §3.) Partial integration shows that if $(\log g)' = 2r$ and $f'(0+) = 0$ then $\int_0^\infty \mathcal{L}f(x) g(x) dx = 0$ which means that

$$g(z) = c \exp(-\beta/8 e^z + 2c_1 z)$$

gives a stationary density. Since $\int_0^\infty \chi(z)g(z) dz = \mathbf{E} \chi(Z(0)) < \infty$, the convergence (48) follows. This shows that

$$\kappa_\beta = \mathbf{E} q_\infty = \mathbf{E} \left[e^{\psi(\tilde{Y}_\infty)} p_1(\tilde{Y}_\infty(0)) \right] < \infty.$$

The only thing left to prove is that $\kappa_\beta = \mathbf{E} q_\infty$ is not zero. The definitions of q and ψ yield

$$q_\infty \geq c p_1(\tilde{Y}_\infty(0)) e^{(2-\beta/2)Y_\infty(0)^+}.$$

By Lemma 5 the function $p_1(\cdot)$ is positive. Since $\tilde{Y}_\infty(0)$ is a.s. finite and we get that $\mathbf{E} q_\infty > 0$ which completes the proof of Theorem 1. \square

Acknowledgments. This research is supported by the Sloan and Connaught grants, the NSERC discovery grant program, and the Canada Research Chair program (Virág). Valkó is partially supported by the Hungarian Scientific Research Fund grant K60708.

We are grateful to Peter Forrester for introducing us to the original work of Dyson (1962), which was the starting point of this paper. We thank Laure Dumaz, Mu Cai and Yang Chen for helpful comments on a previous version.

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